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Original article

Use of hydroxychloroquine in hospitalised COVID-19 patients is associated with reduced mortality: Findings from the observational multicentre Italian CORIST study

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ABSTRACT

Background: Hydroxychloroquine (HCQ) was proposed as potential treatment for COVID-19.

Objective: We set-up a multicenter Italian collaboration to investigate the relationship between HCQ therapy and COVID-19 in-hospital mortality.

Methods: In a retrospective observational study, 3,451 unselected patients hospitalized in 33 clinical centers in Italy, from February 19, 2020 to May 23, 2020, with laboratory-confirmed SARS-CoV-2 infection, were analyzed. The primary end-point in a time-to event analysis was in-hospital death, comparing patients who received HCQ with patients who did not. We used multivariable Cox proportional-hazards regression models with inverse probability for treatment weighting by propensity scores, with the addition of subgroup analyses.

Results: Out of 3,451 COVID-19 patients, 76.3% received HCQ. Death rates (per 1,000 person-days) for patients receiving or not HCQ were 8.9 and 15.7, respectively. After adjustment for propensity scores, we found 30% lower risk of death in patients receiving HCQ (HR=0.70; 95%CI: 0.59 to 0.84; E-value=1.67). Secondary analyses yielded similar results. The inverse association of HCQ with inpatient mortality was particularly evident in patients having elevated C-reactive protein at entry.

Conclusions: HCQ use was associated with a 30% lower risk of death in COVID-19 hospitalized patients. Within the limits of an observational study and awaiting results from randomized controlled trials, these data do not discourage the use of HCQ in inpatients with COVID-19.

1 1. Introduction

The aminoquinoline hydroxychloroquine (HCQ) has been exten-2 sively used in the treatment of malaria and is currently widely used 3 to treat autoimmune diseases like rheumatoid arthritis (RA), systemic 4 lupus erythematosus (SLE) and anti-phospholipid syndrome (APS), due 5 to its immunomodulatory and anti-thrombotic properties [1]. More re-6 cently, a promising role of HCQ has been suggested in viral infec-7 tions [2], since it directly inhibits viral entry and spread in several in 8 vitro and in vivo models. Due to these properties, HCQ has been used 9 in Ebola virus disease [3,4], human immunodeficiency virus (HIV) in-10 11 fection [5], SARS-CoV-1 infection and the Middle East Respiratory Syndrome (MERS) [6,7] and gained worldwide attention as a possible ther-12 apy in COVID-19 patients [8]. 13

14 HCQ might inhibit the intracellular glycosylation of ACE 2, the re-15 ceptor used by the SARS-CoV-2 virus to enter the cells, resulting in a reduced ligand recognition and internalization of the virus [7] and exert-16 ing a possible protective role in SARS-CoV-2 infection. Moreover, due to 17 its immunomodulatory, anti-inflammatory and anti-thrombotic effects, 18 HCQ could also modulate the severity of the disease. However, the exact 19 mechanism for the potential benefit in COVID-19 is largely speculative 20 [9] and might be counterbalanced by adverse effects, mainly cardiovas-21 cular [10,11], so that the net balance of this drug's use remains to be 22 established. 23

The American Food and Drug Administration (FDA) allowed Chloro-24 quine (CQ) phosphate and HCQ to be provided to certain hospitalized 25 patients because these drugs may possibly help patients with severe 26 COVID-19 [12]. The European Medicines Agency (EMA) authorized the 27 use of CQ and HCQ for COVID-19 in clinical trials or as emergency use 28 [13], while the Italian Drug Agency (AIFA) stated in this emergency 29 phase that therapeutic use of HCQ might be considered in COVID-19 30 patients, both in those with mild presentation managed at home and in 31

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hospitalized patients [14]. In clinical practice, HCQ rather than chloroquine has been used because of its more potent antiviral properties and
better safety profile [15].

However, in the light of a recent publication [16], that was later retracted [17], on the lack of safety and efficacy of HCQ in the treatment for COVID-19 patients the Executive Group of the Solidarity Trial decided to implement a temporary pause of the HCQ arm within the trial as a precaution, while the safety data is being reviewed [18]. Similarly, the Italian drug Agency AIFA decided to suspend the authorization to use HCQ for COVID-19 treatment outside clinical trials [19].

42 Recent reviews of clinical trials or observational studies [20-43 24] have reported insufficient and often conflicting evidence on the benefits and harms of using HCQ to treat COVID-19 and concluded that as 44 45 such, it was impossible to determine the balance of benefits to harm. Until now, although several trials had been started on the use of CQ and 46 HCQ in COVID-19, only few of them have been published [25] on small 47 numbers of patients or on surrogate endpoints or in exposed subjects for 48 prophylaxis use [26]. 49

While waiting the results from ongoing randomized clinical trials 50 (RCT) to define the efficacy in preventing hard endpoints of this treat-51 ment so widely used during the emergency phase of the COVID-19 pan-52 demic, powered retrospective observational studies performed in differ-53 54 ent geographical and disease conditions may still be useful to shed light 55 on this debate. Two retrospective observational studies, both conducted in the New York metropolitan region, did not report any significant as-56 sociation between HCQ use and rates of intubation or death [27,28]. 57

No data are presently available from large cohorts of patients in Italy, 58 59 which represents one of the most affected countries in terms of total deaths for COVID-19 in the world [29]. We undertook a multicenter Ital-60 ian collaboration [30] to investigate the relationship between underly-61 ing risk factors and COVID-19 outcomes, and to evaluate the association 62 63 between different drug therapy and disease severity and/or mortality. We report here the results obtained in 3,451 hospitalized COVID -19 64 patients receiving or not HCQ treatment. 65

66 2. Material and methods

67 2.1. Setting

This national retrospective observational study was conceived, co-68 ordinated and analysed within the CORIST Project (ClinicalTrials.gov 69 ID: NCT04318418, 30]. The study was approved by the institutional 70 ethics board of the Istituto di Ricovero e Cura a Carattere Scientifico 71 72 (IRCCS) Neuromed, Pozzilli, and of all recruiting centres. Data for the 73 present analyses were provided by 33 hospitals distributed throughout Italy (listed in the supplementary file). Acceptance to participate in the 74 project or to provide data for the present analysis was not related to 75 the use of CQ/HCQ. Each hospital provided data from hospitalized pa-76 77 tients who had a positive test result for the SARS-CoV-2 virus at any time during their hospitalization from February 19 to May 23, 2020. 78 The follow-up continued through May 29, 2020. 79

80 2.2. Data sources

We developed a cohort comprising 3,971 patients with laboratory-81 confirmed SARS-CoV-2 infection in an in-patient setting. The SARS-CoV-82 83 2 status was declared based on laboratory results (polymerase chain reaction on nasopharyngeal swab) from each participating hospital. Clin-84 85 ical data were abstracted at one-time point from electronic medical records or charts, and were collected using either a centrally designed 86 electronic worksheet or a centralized web-based database. Collected 87 data included patients' demographics, laboratory test results, medica-88 tion administration, historical and current medication lists, historical 89 and current diagnoses, and clinical notes. In addition, specific infor-90 mation on the most severe manifestation of COVID-19 occurred during 91

hospitalization was retrospectively captured. Maximum clinical sever-92 ity observed was classified as mild pneumonia; or severe pneumonia; or 93 acute respiratory distress syndrome (ARDS) [31]. Specifically, we ob-94 tained the following information for each patient: hospital; date of ad-95 mission and date of discharge or death; age; sex; the first recorded inpa-96 tient laboratory tests at the entry (creatinine, C-reactive protein); past 97 and current diagnoses (myocardial infarction, heart failure, diabetes, 98 hypertension, respiratory disease and cancer) and current drug thera-99 pies for COVID-19 - HCQ, lopinavir/ritonavir or darunavir/cobicistat, 100 remdesevir, tocilizumab or sarilumab, corticosteroids, heparin, and for 101 comorbidities (insulin, anti-hypertensive treatments, aldosterone recep-102 tor antagonists, diuretics, statins, sacubitril/valsartan). A diagnosis of 103 pre-existing cardiovascular disease was based on history of myocar-104 dial infarction or heart failure. Chronic kidney disease was classified 105 as: stage 1: kidney damage with normal or increased glomerular fil-106 tration rate (GFR) (>90 mL/min/1.73 m²); stage 2: mild reduction 107 in GFR (60-89 mL/min/1.73 m²); stage 3a: moderate reduction in 108 GFR (45-59 mL/min/1.73 m²); stage 3b: moderate reduction in GFR 109 (30-44 mL/min/1.73 m²); stage 4: severe reduction in GFR (15-29 110 mL/min/1.73 m²); stage 5: kidney failure (GFR <15 mL/min/1.73 m²) 111 or dialysis). For statistical analysis, stages 3a and 3b and stages 4 and 112 5 were combined. GFR was calculated by the Chronic Kidney Disease 113 Epidemiology Collaboration (CKD-Epi) equation. Patients were defined 114 as receiving HCQ if they were receiving it at admission to hospital or 115 received it during the follow-up period. According to the AIFA guidance 116 [14], HCQ was administered at dose of 400 mg x 2/day or x4/day the 117 first day, and 200 mg x 2/day from the second day onwards for at least 118 5 to a maximum of 10 days, according to the clinical evolution of the 119 disease. 120

2.3. Statistical analysis

The study index date was defined as the date of hospital admission. 122 Index dates ranged from February 19, 2020 to May 23, 2020. The study 123 end point was the time from study index to death. The number of pa-124 tients who either died, or had been discharged alive, or were still admit-125 ted to hospital as of May 29, 2020, were recorded, and hospital length of 126 stay was determined. Patients alive had their data censored on the date 127 of discharge or as the date of the respective clinical data collection. Data 128 were censored at 35 days of follow up in n=330 (8.3%) patients with a 129 follow up greater than 35 days. 130

121

Of the initial cohort of 3,971 patients, 350 patients were excluded 131 from the analysis because they had at least one missing data at baseline 132 or lost to follow up on HCQ use (N=94), other drug therapies for COVID-133 19 (n=265), time to event (n=59), outcome (death/alive, n=8), COVID-134 19 severity (n=4), age (n=4 with missing data and n=2 with age<18 135 years) or sex (n=2). Of the remaining 3,621 patients, 170 patients died 136 or were discharged within 24 hours after presentation, and were also 137 excluded from the analysis. 138

At the end, the analysed cohort consisted of n=3,451 patients. In pa-139 tients not included in the analysis (n=520), as unique difference with the 140 analysed group, the prevalence of diabetics (19.9% vs 14.8%, P=0.0066) 141 and, to a less extent, of men (62.3% vs 58.3%, P=0.081) was higher. Out 142 of 3,541 patients, 295 (8.5%) had at least a missing value for covariates. 143 Distribution of missing values was as follows: n=178 for C-reactive pro-144 tein; n=69 for GFR; n=74 for history of ischemic disease; n=64 for his-145 tory of chronic pulmonary disease; n=51 for diabetes; n=51 for hyper-146 tension and n=56 for cancer. We used multiple imputation techniques 147 (SAS PROC MI, n=10 imputed datasets; and PROC MIANALYZE) to max-148 imize data availability. As sensitivity analysis, we also conducted a case-149 complete analysis on 3,156 patients. 150

Cox proportional-hazards regression models were used to estimate 151 the association between HCQ use and death. Since multiple imputation 152 was applied, the final standard error was obtained using the Rubin's 153 rule based on the robust variance estimator in Cox regression [32]. The 154 proportional hazards assumption was assessed using weighed Schoen-

feld residuals, and no violation was identified. To account for the non-156 randomized HCQ administration and to reduce the effects of confound-157 158 ing, the propensity-score method was used. The individual propensities for receiving HCQ treatment were assessed with the use of a multivari-159 able logistic-regression model that included age, sex, diabetes, hyperten-160 sion, history of ischemic heart disease, chronic pulmonary disease, GFR, 161 C-reactive protein, hospitals clustering and use of other drug therapies 162 for COVID-19 (lopinavir/ritonavir or darunavir/cobicistat, remdesivir, 163 164 corticosteroids, tocilizumab or sarilumab). Associations between HCQ treatment and death was then appraised by multivariable Cox regres-165 166 sion models with the use of propensity-score and further controlling for hospitals clustering as random effect (frailty model). The use of a 167 frailty model was chosen as suggested in [33]. The primary analysis 168 169 used inverse probability by treatment weighting; the predicted probabilities from the propensity-score model was used to calculate the sta-170 bilized inverse-probability-weighting weight [34]. Stabilized weights 171 were normalized so that they added up the actual sample size. Sec-172 ondary analyses used propensity-score stratification (n=5 strata) or mul-173 tivariable Cox regression analysis or multivariable logistic regression 174 analyses comparing death versus alive patients, or accounted for hos-175 pitals clustering via stratification or by robust sandwich estimator. Pre-176 established subgroup analyses were conducted according to age or sex 177 178 of patients, degree of COVID-19 severity experienced during the hospi-179 tal stay, C-reactive protein at basal or other drug therapies for COVID-19. Hospitals were clustered according to their geographical distribu-180 tion, as illustrated in Table 1. To quantify the potential for an un-181 measured confounder to render apparent statistically significant hazard 182 183 ratio non-significant, the E-value was calculated [35]. Analyses were performed with the aid of the SAS version 9.4 statistical software for 184 Windows. 185

186 3. Results

187 We included in the final current analyses 3,451 patients who were hospitalized with confirmed SARS-CoV-2 infection at 33 clinical centres 188 189 across Italy and either died, had been discharged, or were still in hospi-190 tal as of May 29, 2020. Of these patients, 2,634 (76.3%, range among hospitals 53.2% to 93.6%) received HCQ. Timing of the first dose of 191 192 HCQ after presentation to the hospital was 1 day for the large majority of centres, and 2 to 3 days for the others. HCQ was administered in all 193 centres at the dose of 400 mg/day (in one centre however it was used at 194 the dose of 600 mg/day and in another at the dose of 600 mg/day but 195 only in patients younger than 65 years). Duration of treatment ranged 196 from 5 to 15 days (with 10 days as the modal value). The drug used was 197 198 HCO in all hospitals.

Baseline characteristics according to HCO use are shown in Table 1. 199 Patients receiving HCQ were more likely younger, men and had higher 200 levels of C-reactive protein and less likely had ischemic heart disease, 201 202 cancer or stages 3a or greater chronic kidney disease (Table 1). Patients receiving HCQ more likely received another drug for COVID-19 treat-203 ment (78.4%; lopinavir/ritonavir or darunavir/cobicistat, remdesevir, 204 205 tocilizumab or sarilumab, corticosteroids), in comparison with non-HCQ patients (46.3%; P<0.0001; Table 1). 206

The unadjusted differences and differences adjusted by propensity scores between HCQ-treated and non-HCQ treated patients for each variable included in the propensity score are shown in Fig. 1. All the pretreatment differences disappeared after adjustment by propensity score weighting. The C-statistic of the propensity-score model was 0.74.

212 3.1. Primary outcome

Out of 3,628 patients, 576 died (16.7%), 2,390 were discharged alive (69.3%) and 485 (14.1%) were still at the hospital. The median followup was 14 days (interquartile range 8 to 22; range 2 to 35; 55,388 person-days). Death rate (per 1,000 person-days) was 8.9 in HCQ and 238

15.7 in non-HCQ patients (Table 2). At univariable analysis, hazard ra-217 tio for mortality was 0.56 (95%CI: 0.47 to 0.67). In the primary mul-218 tivariable analysis with inverse probability weighting according to the 219 propensity score, HCQ use was associated with a 30% (95%CI: 16% to 220 41%) reduction in death risk (Fig. 2, Table 2, E-value=1.67). Secondary 221 multivariable analyses yielded very similar results (Table 2), as well as 222 case-complete analyses restricted to the 3,156 patients without missing 223 data (Table 2). Considering secondary multivariable analyses overall, 224 HR for mortality associated with HCQ ranged between 0.64 to 0.70, 225 according to type of analyses. Control of hospitals clustering with dif-226 ferent approaches also yielded similar results for the primary analysis 227 (HR=0.71, 95%CI: 0.59 to 0.85 when hospitals clustering was stratified 228 for and HR=0.69, 95%CI: 0.54 to 0.88 with the robust sandwich esti-229 mator). 230

Subgroup analyses are presented in **Table 3**. HCQ use remained consistently associated with reduced mortality in almost all subgroups. The inverse association of HCQ with inpatient mortality is slightly more evident in women, elderly and in patients who experienced a higher degree of COVID-19 severity. It was absent in-patient with C-reactive protein <10 mg/L and clearly confined to patients with elevated C-reactive protein (**Table 3**). 237

4. Discussion

In a large cohort of 3,451 patients hospitalized for COVID-19 in 33 239 clinical centers all over Italy, covering almost completely the period of 240 the hospitalization for COVID-19, the use of HCQ was associated with 241 a significant better survival. In-hospital crude death rate was 8.9 per 242 1,000 person-day for patients receiving HCQ and 15.7 for those who 243 did not. After adjustment for known possible confounders, we observed 244 a 30% reduction in the risk of death in patients receiving HCQ therapy 245 as compared with those who did not. 246

Our findings provide clinical evidence in support of guidelines by 247 Italian and several international Societies suggesting to use HCQ ther-248 apy in patients with COVID-19. However, the observed associations 249 should be considered with caution, as the observational design of our 250 study does not allow to fully excluding the possibility of residual con-251 founders. Large randomized clinical trials in well-defined geographical 252 and socio-economic conditions and in well-characterized COVID-19 pa-253 tients, should evaluate the role of HCQ before any firm conclusion can 254 be reached regarding a potential benefit of this drug in patients with 255 COVID-19. 256

Over 76% of patients received HCQ either alone or in combina-257 tion with other drugs. They were more likely to be younger, men 258 and with higher levels of C reactive protein at entry, while less likely 259 had pre-existing comorbidities such as ischemic heart disease, cancer 260 and severe chronic kidney disease, as compared to patients not receiv-261 ing the drug. We adjusted our analyses for possible confounders, in-262 cluding age, sex, diabetes, hypertension, history of ischemic heart dis-263 ease, chronic pulmonary disease, chronic kidney disease, C-reactive pro-264 tein and additional treatments for COVID-19, and took into account 265 possible differences across centres by either adjustment or stratifica-266 tion. To minimize bias due to the observational design, we used dif-267 ferent analytical approaches aiming at creating an overall balance be-268 tween comparison groups. Finally, we tried to limit bias due to miss-269 ing data by using a multiple imputation approach, but in no case, 270 the result was changed. Despite all these precautions, we recognize 271 the possibility, however, of residual unmeasured confounders affecting 272 results. 273

Systematic reviews of small clinical trials had reported contrasting results that were however scarcely reliable because of poor designs [20–25]. The HCQ doses tested in a Chinese randomized clinical trial [25] were approximately double as compared to that used in our study (1200 mg vs 800 mg as loading dose, 800 mg vs 400 mg as maintenance dose) for twice the time (14-21 days versus 7-10 days). National guidelines in Italy suggest to use HCQ 200 mg

Table 1

General characteristics of COVID-19 patients at baseline, according to hydroxychloroquine use.

	Hydroxychloroquine				
Characteristic	No (N=817)	Yes (N=2,634)	P-value unadjusted*		
Age-median (IQR-yr.)	73 (58-83)	66 (55-77)	<.0001		
Gender- no (%)			<.0001		
Women	361 (44.2%)	940 (36.7%)			
Men	456 (55.8%)	1,694 (64.3%)			
Diabetes- no (%)			0.71		
No	633 (77.5%)	2,090 (79.3%)			
Yes	162 (19.9%)	515 (19.6%)			
missing data	22 (2.7%)	29 (1.1%)			
Hypertension - no (%)			0.31		
No	378 (46.3%)	1,294 (49.1%)			
Yes	416 (50.9%)	1,312 (49.8%)			
missing data	23 (2.7%)	28 (1.1%)			
Ischemic heart disease- no (%)			<.0001		
No	610 (74.7%)	2,190 (83.1%)			
Yes	179 (21.9%)	398 (15.1%)			
missing data	28 (3.4%)	46 (1.8%)			
Chronic pulmonary disease- no (%)		. ,	0.21		
No	666 (81.5%)	2,225 (84.5%)			
Yes	127 (15.5%)	369 (14.0%)			
missing data	24 (2.9%)	40 (1.5%)			
Cancer- no (%)	21 (210/0)	10 (110,0)	0.036		
No	694 (84.9%)	2,338 (88.8%)	0.030		
Yes	101 (12.4%)	262 (9.9%)			
missing data	22 (2.6%)	34 (1.3%)			
CKD stage**- no (%)	22 (2.0%)	54 (1.5%)	<.0001		
Stage 1	241 (29.5%)	970 (36.8%)	<.0001		
8					
Stage 2 Stage 2a or stage 2b	281 (34.4%)	991 (37.6%)			
Stage 3a or stage 3b	180 (22.0%)	487 (18.5%)			
Stage 4 or stage 5	89 (10.9%)	143 (5.4%)			
missing data	26 (3.2%)	43 (1.6%)	0.0000		
C Reactive Protein- no (%)			0.0003		
<1 mg/L	104 (12.7%)	256 (9.7%)			
1-3 mg/L	120 (14.7%)	301 (11.4%)			
>3 mg/L	549 (67.2%)	1,943 (73.8%)			
missing data	44 (5.4%)	134 (5.1%)			
Lopinavir or Darunavir use			<.0001		
No	621 (76.0%)	1,203 (36.7%)			
Yes	196 (24.0%)	1,431 (64.3%)			
Tocilizumab or Sarilumab use			<.0001		
No	755 (92.4%)	2,160 (82.0%)			
Yes	62 (7.6%)	474 (18.0%)			
Remdesivir use			0.0015		
No	808 (98.9%)	2,551 (96.9%)			
Yes	9 (1.1%)	83 (3.1%)			
Corticosteroids use			<.0001		
No	596 (73.0%)	1,655 (62.8%)			
Yes	221 (27.0%)	979 (37.2%)			
Clusters of hospitals	. ,		<.0001		
Northern regions (except Milan) (n)	169 (20.7%)	616 (23.4%)			
Milan (m)	161 (19.7%)	525 (19.9%)			
Center regions (except Rome) (c))	303 (37.1%)	747 (28.4%)			
Rome (r)	94 (11.5%)	390 (14.8%)			
Southern regions (s)	90 (11.0%)	356 (13.5%)			
Southern regions (s)	50 (11.0%)	JJU (13.5%)			

(n) include hospitals of Novara, Monza, Varese, Pavia, Cremona and Padova; (m) include Humanitas Clinical and Research Hospital, Centro Cardiologico Monzino, and hospitals of San Donato Milanese (Milano) and Cinisello Balsamo (Milano); (c) include hospitals of Modena, Ravenna, Forlì, Firenze, Pisa, Chieti and Pescara; (r) include National Institute for Infectious Diseases "L. Spallanzani" and Università Cattolica del Sacro Cuore; (s) include hospital of Napoli, Pozzilli (Isernia), Acquaviva delle Fonti (Bari), Foggia, Taranto, Catanzaro, Catania and Palermo *Chi-square test. **Stage 1: Kidney damage with normal or increased glomerular filtration rate (GFR) (>90 mL/min/1.73 m²); Stage 2: Mild reduction in GFR (60-89 mL/min/1.73 m²); Stage 3a: Moderate reduction in GFR (45-59 mL/min/1.73 m²); Stage 3b: Moderate reduction in GFR (30-44 mL/min/1.73 m²); Stage 4: Severe reduction in GFR (15-29 mL/min/1.73 m²); Stage 5: Kidney failure (GFR < 15 mL/min/1.73 m² or dialysis).

twice daily for at least 5-7 days in patients over 70 years and/or with co-morbidities (chronic obstructive pulmonary disease, diabetes, cardiovascular disease) even with mild respiratory symptoms or with radiographically documented pneumonia or in severe patients [36]. The lower doses of HCQ used in our centers, as suggested by Italian official guidelines [19,36], may have been both more effective and 286 safer. 287

Two recently published large observational studies, both from large hospitals in New York City, showed no association between HCQ use and in-hospital mortality [27,28], and deserve specific discussion. In the 290

Table 2

Incidence rates and hazard ratios for death in COVID-19 patients, according to hydroxychloroquine use.

Multiple imputation an	alysis (N=3,451)			
	Death (N=576)	Patient at risk (N=3,451)	Person-days	Death Rate (x1,000 person-days)
Hydroxychloroquine				
No- no. (%)	190 (23.3%)	817 (100%)	12,084	15.7
Yes- no. (%)	386 (14.7%)	2,634 (100%)	43,304	8.9
Hazard ratio for death (HCQ versus non HCQ)			HR (95% CI)	
Crude analysis			0.56 (0.47 to 0.67)	
Multivariable analysis*			0.70 (0.58 to 0.85)	
Propensity score analysis, inverse probability weighting** (primary analysis)			0.70 (0.59 to 0.84)	
Propensity score analysis, stratification (n=5 strata)**			0.67 (0.56 to 0.81)	
Odds ratio for death (HCQ versus non HCQ)			OR (95% CI)	
Propensity score analysis, inverse probability weighting**			0.67 (0.54 to 0.82)	
Case Complete Analy	ysis (N=3,156)			
	Death (N=510)	Patient at risk (N= 3,156)	Person-days	Death Rate (x1,000 person-days)
Hydroxychloroquine	!			
No- no. (%)	170 (22.9%)	741 (100%)	11,050	15.4
Yes- no. (%)	340 (14.1%)	2,415 (100%)	39,274	8.7
Hazard ratio for dea	Hazard ratio for death (HCQ versus non HCQ)			HR (95% CI)
Crude analysis			0.56 (0.46 to 0.67)	
Multivariable analysis*			0.71 (0.59 to 0.86)	
Propensity score analysis, inverse probability weighting**			0.64 (0.53 to 0.76)	
Propensity score analysis, stratification (n=5 strata)**			0.68 (0.56 to 0.82)	
Odds ratio for death (HCQ versus non HCQ)			OR (95% CI)	
Propensity score analysis, inverse probability weighting**				0.67 (0.54 to 0.82)

Abbreviations: HR, hazard ratios; CI, confidence intervals. *Controlling for age, sex, diabetes, hypertension, history of ischemic heart disease, chronic pulmonary disease, chronic kidney disease, C-reactive protein, lopinavir/ritonavir or darunavir/cobicistat, tocilizumab or sarilumab, remdesivir or corticosteroids use as fixed effects and hospitals clustering as random effect. **Including hospitals clustering as random effect covariate.

study of Geleris et al. [27], the percentage use of HCQ was lower than in

²⁹² Italy; moreover, in both US studies [27,28] the drug was more frequently

administered to patients with previous illnesses and a more severe presentation of the disease. Our cohort included milder pneumonia patients

than the US population, due to between-country differences in indica-

tions to the drug for the beginning of therapy (e.g., mild pneumonia in296Italy versus only severe pneumonia and ARDS in the US). Concomitant297use of other drugs for COVID-19 was very low in one study [27] and was298not reported in the other study [28]. In our cohort, patients receiving299HCQ were more likely treated with another drug for COVID-19 treat-300

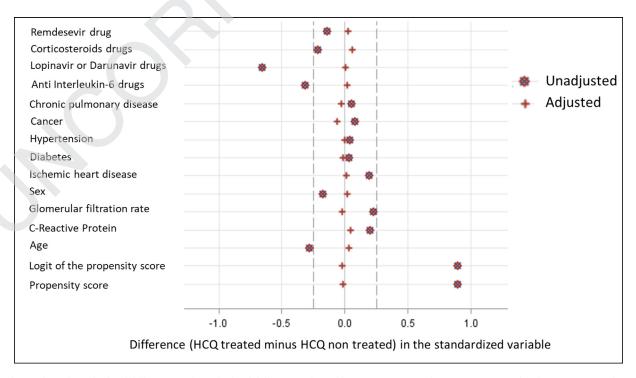


Fig. 1. The unadjusted standardized differences and standardized differences adjusted by propensity scores between HCQ-treated and non-HCQ treated patients for the variables included in the propensity score. All differences for the matched observations are within the recommended limits of –0.25 and 0.25, which are indicated by reference lines.

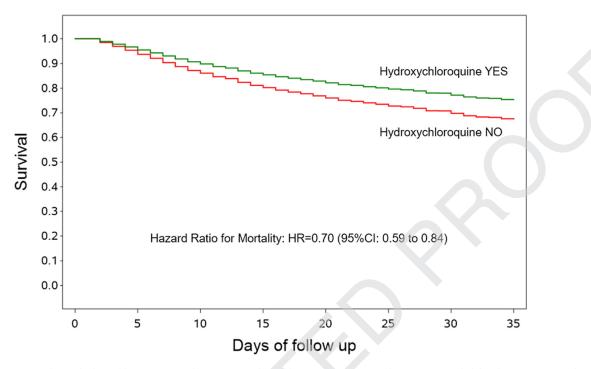


Fig. 2. Survival curves according to hydroxychloroquine use. The curves are adjusted by propensity score analysis (inverse probability for treatment weighting) and hospital index as random effect, and are generated using the first imputed dataset. The other imputed datasets are similar and thus omitted.

Table 3

Hazard ratios for mortality according to hydroxychloroquine use in different subgroups.

	Hydroxychloroquine NO (N=817)	Hydroxychloroquine YES (N=2,634)				
Subgroups	No. death/patient at risk	No. death/patient at risk	HR (95% CI)*			
Women	80/361	116/940	0.63 (0.46 to 0.86)			
Men	110/456	270/1,694	0.74 (0.60 to 0.93)			
Age <70 years	22/357	93/1,542	0.76 (0.50 to 1.16)			
Age ≥70 years	168/460	293/1,092	0.68 (0.56 to 0.83)			
Highest degree of COVID-19 severity experienced at hospital						
Mild pneumonia or less	28/424	40/1,358	0.70 (0.41 to 1.18)			
Severe pneumonia	80/253	172/764	0.76 (0.58 to 0.99)			
Acute respiratory distress syndrome	82/140	174/512	0.68 (0.52 to 0.90)			
Use of other COVID-19 treatments						
No	101/439	64/570	0.63 (0.45 to 0.88)			
Yes	89/378	322/2,064	0.77 (0.61 to 0.99)			
C-Reactive Protein at basal**						
<10 mg/L	56/412	125/1,138	1.23 (0.86 to 1.77)			
≥10 mg/L	123/361	241/1,362	0.59 (0.47 to 0.73)			

Abbreviations: HR, hazard ratios; CI, confidence intervals; *Propensity score analysis, inverse probability weighting, including hospital clustering as random effect covariate; multiple imputed analysis.

Lopinavir/ritonavir or darunavir/cobicistat or tocilizumab or sarilumab or remdesivir or corticosteroids.

**Missing data for N=178. Frequencies and hazard ratios are based on a case complete analysis (N=3,273) without missing data for C-reactive Protein; multiple imputed analysis (N=3,451) yielded very similar results.

ment (78.4%), in comparison with non-HCQ patients (46.3%). Anyway,
our findings are adjusted for concomitant other drugs use.

While the US studies were confined to one hospital only or a defined 303 relatively small area in the Country, our study included 33 hospitals 304 distributed all over Italy, covering regions with a high number of cases 305 and a high intra-hospital mortality and regions with a lower burden 306 of the disease. The participating Italian clinical centers have different 307 healthcare facilities, different size, specialization, and ownership, and 308 therefore quite closely represent the real-life Italian approach to COVID-309 310 19. Moreover, they differed for the percentage of use of HCQ and for the rate of in-hospital mortality that ranged between 34.1 and 1.5 per 311 1,000 persons/day. To consider this variability, we adjusted the analy-312 sis for recruiting center and performed a number of subgroup analyses. 313 314 In all circumstances, the association between HCQ use and a reduced 315 risk of death of about 30% was maintained. Quite interestingly, the inverse association of HCQ with inpatient mortality was more evident in 316 elderly, in patients who experienced a higher degree of COVID-19 sever-317 ity or especially having elevated C-reactive protein, suggesting that the 318 anti-inflammatory potential of HCQ may have had more important role 319 rather than its antiviral properties. HCQ, indeed, beside an antiviral ac-320 tivity, may have both anti-inflammatory and anti-thrombotic effects [8]. 321 This can justify its effect in reducing mortality risk, since Sars-Cov-2 can 322 induce pulmonary microthrombi and coagulopathy, that are a possible 323 cause of its severity [37,38] and the lack in preventing SARS-CoV-2 in-324 fection after exposure [26] 325

Nevertheless, large randomized clinical trials on the efficacy of HCQ 326 on hard end-points are still lacking and the largest observational study 327 showing no effect in reducing mortality has been retracted [16,17], 328 Agencies have suspended clinical trials on the efficacy of HCQ on 329 COVID-19 disease or have restricted its use only to patients included 330

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in clinical trials, in the absence of an ample, serene and balanced discussion at international level.

Very recently, a large RCT has become available as a pre-print publication [39], reporting no beneficial effect of HCQ in patients hospitalized with COVID-19. However, the dose of HCQ used in that trial was almost the double of that administered in our real life conditions. A reduced mortality was also observed by other observational studies using low or intermediate doses of HCQ [40,41].

Moreover, in our study patients taking HCQ more frequently received other anti-COVID drugs, whose interaction in reducing mortality cannot be completely ruled-out. Of note, despite the higher dosage used, the RCT did not show any excess in ventricular tachycardia or ventricular fibrillation in the HCQ arm (39).

Therefore, it will be very important to compare results of studies with different mode of use and doses of HCQ, different characteristics of treated and untreated patients and different academic or real-world conditions.

348 4.1. Strengths and limitations

A major strength of this study is the large, unselected patient sample from 33 hospitals, covering the entire Italian territory. Patient sampling covered all the overt epidemic period in Italy. Several statistical approaches were used to overcome biases due to the observational nature of the investigation.

This study has however, several recognized limitations. The study 354 355 population pertains to Italy, and the results obtained may not be appli-356 cable to other populations with a possibly different geographical and socio-economic conditions and natural history of COVID-19. Due to the 357 retrospective nature of the study, some parameters were not available 358 in all patients, and all in-hospital medications might have been not fully 359 360 recorded. Moreover, although guidelines on the use of HCQ in COVID-19 patients had been published in Italy since the first phase of the pan-361 362 demic, individual centers could have deviated from recommendations and used different doses or treatment schemes. We have no information 363 on the HCQ doses used individually nor of their possible association with 364 azithromycin. Moreover, adverse events possibly related to drug therapy 365 366 were not collected, thus we cannot exclude bias due to therapy interruption because of side effects; we do not know whether some deaths 367 could have been due to cardiovascular complications of HCO. However, 368 recent data on Italian wards showed that COVID-19 patients receiving 369 HCQ and azithromycin had a QTc-interval longer than before therapy, 370 but did not experience, during their hospital stay, any arrhythmic com-371 plications, such as syncope or life-threatening ventricular arrhythmias 372 [42], a finding also reported by the RCT mentioned above (39). 373

Finally, the possibility of unmeasured residual confounding cannot be completely ruled-out. However, the E-value for the lower boundary of the confidence interval of our main result is 1.67, indicating that the confidence interval could be moved to include the null by a strong unmeasured confounder associated with both HCQ treatment and death with a risk ratio of 1.67-fold for each, above and beyond all the measured confounders. Weaker confounders, however, could not do so.

381 5. Conclusions

Our study, including a large real life sample of patients hospitalized with COVID-19 all over Italy, shows that HCQ use (200 mg twice/day) was associated with a 30% reduction of overall in-hospital mortality. In the absence of clear-cut results from controlled, randomized clinical trials, our data do not discourage the use of HCQ in inpatients with COVID-19. Given the observational design of our study, however, these results should be transferred with caution to clinical practice.

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Declaration of Competing Interests

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The Authors alone are responsible for the views expressed in this Article. They do not necessarily represent the views, decisions, or policies of the Institutions with which they are affiliated.

Appendix 1

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Supplementary materials 626

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